### Coupled length scales in eroding landscapes

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We report results from an empirical study of the anisotropic structure of eroding landscapes. By constructing a novel correlation function, we show quantitatively that small-scale channel-like features of landscapes are coupled to the large-scale structure of drainage basins. We show additionally that this two-scale interaction is scale-dependent. The latter observation suggests that a commonly applied effective equation for erosive transport may itself depend on scale.

DOI: 10.1103/PhysRevE.63.055102

PACS number(s): 05.65.+b, 92.40.Fb, 92.40.Gc, 45.70.Qj

Landscapes erode in part due to the shearing stresses imposed by the downhill transport of water, sediment, and other material [1]. Whereas the effects of erosion appear nearly obvious to the eye, they are notoriously hard to quantify. For example, the functional form of many river-network scaling laws [2] may be obtained from simple graphical constructions [3,4] that have no obvious relation to real surfaces, eroded or not. Moreover, the appropriate partial differential equations for dynamic modeling are a source of much controversy [5-21]. Here we report results from analyses of eroded topography that explicitly quantify a unique aspect of landscape erosion: a coupling between small-scale channellike features and the large-scale structure of drainage basins. A detailed study of this two-scale interaction reveals a rich hierarchy of scale dependencies in erosive processes. We provide evidence that these same scale dependencies are implicitly present in a commonly applied phenomenological theory of erosive transport [5-16].

We probe the detailed anisotropic correlations of a landscape  $h(\mathbf{r})$ , where *h* is elevation and **r** is position. Our analysis is inspired by recent theoretical results [19,20] that point out the role of anisotropy in erosion and by recent empirical analyses of other pattern forming systems [22,23]. We first form the function

$$w(\mathbf{x}, \boldsymbol{\phi}; L_c) = e^{i\mathbf{k}_{\boldsymbol{\phi}} \cdot \mathbf{x}} g(\mathbf{x}; L_c). \tag{1}$$

The function *w* is built from a modulated plane wave. The plane wave's angular orientation  $\phi$  and wavelength  $\lambda_0$  determine the wave vector

$$\mathbf{k}_{\phi} = \frac{2\pi}{\lambda_0} (\cos\phi, \sin\phi), \quad 0 \le \phi < 2\pi, \tag{2}$$

and the plane wave is modulated by a Gaussian taper g with standard deviation  $L_c/4$ :

$$g(\mathbf{x};L_c) = L_c^{-2} e^{-8|\mathbf{x}|^2/L_c^2}.$$
 (3)

We correlate w with the local topography over a region of size  $L_c$  by computing

$$R(\mathbf{r},\phi;L_c) = \left| \int_{|\mathbf{x}-\mathbf{r}| \leq L_c/2} d\mathbf{x} \ w(\mathbf{x}-\mathbf{r},\phi;L_c)h(\mathbf{x}) \right|.$$
(4)

We then define the local *channelization vector*  $\mathbf{c}(\mathbf{r})$  from the phase of the  $\pi$ -periodic component of  $R(\mathbf{r}, \phi)$ , i.e.,

$$\mathbf{c}(\mathbf{r}) = [-\operatorname{Im}(Z), \operatorname{Re}(Z)], \quad Z = \int_0^{2\pi} e^{-2\phi i} R(\mathbf{r}, \phi) d\phi.$$
(5)

When topographic height-height correlations  $\langle |h(\mathbf{x}) - h(0)|^2 \rangle$  are elliptically anisotropic, **c** should be perpendicular to the dominant local wave vector with magnitude  $2 \pi / \lambda_0$ . In other words, when  $\lambda_0$  is of the order of the width of channels,  $\pm \mathbf{c}$  points in the direction in which they flow [24], and |**c**| is proportional to the channel depth. The vector field **c**(**r**) varies over all length scales greater than  $L_c$ .

To complete our description, we define the coarse-grained slope

$$\mathbf{s}(\mathbf{r};L_s) = G^{-1} \int_{|\mathbf{x}-\mathbf{r}| \le L_s/2} d\mathbf{x} \ g(\mathbf{x}-\mathbf{r};L_s) \cdot \boldsymbol{\nabla} h(\mathbf{x}), \quad (6)$$

where the normalization factor  $G^{-1} = \int_{|\mathbf{x}| \leq L_s/2} d\mathbf{x} g(\mathbf{x}; L_s)$ . We calculate the angular difference  $\delta \theta$  between **c** and **s** and study the averaged quantity

$$\langle \cos^2 \delta \theta \rangle = \left\langle \left( \frac{\mathbf{c} \cdot \mathbf{s}}{|\mathbf{c}||\mathbf{s}|} \right)^2 \right\rangle,$$
 (7)

where **c** and **s** are measured at length scales  $L_c$  and  $L_s$ , respectively, and the angle brackets indicate spatial averaging. Equation (7) should be interpreted as the average correlation between a scale-dependent erosive response  $\mathbf{c}(L_c)$  and a scale-dependent stress  $\mathbf{s}(L_s)$  that induces it.

To see how these correlations manifest themselves in real topography, we have studied the United States Geological Survey (USGS) 1-degree digital elevation map that ranges from  $38^{\circ}-39^{\circ}$  north latitude and  $120^{\circ}-121^{\circ}$  west longitude, a rectangular area whose dimensions are roughly 100 km on a side, with a resolution of approximately 90 m. This area, entirely contained within California, ranges from the western flank of the Sierra Nevada mountains to the eastern edge of the Central Valley. A wide range of erosive features exist, including those of glacial, fluvial, and alluvial origin.

Figure 1(a) displays  $\langle \cos^2 \delta \theta \rangle$  as a function of  $L_s/L_c$ , averaged over the entire map. (Here and elsewhere, we study

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FIG. 1. (a) The mean squared cosine of the angular separation  $\delta\theta$  between the slope vector **s** and the channelization vector **c** as a function of the ratio of the respective scales,  $L_s$  and  $L_c$ , at which they are calculated. Here  $L_c = 1.0$  km and the maximum correlation occurs at  $L_s^* \approx 10L_c$ . Curves computed for different values of  $L_c$  are qualitatively similar, but the maximum occurs for a value  $L_s^*$  that depends on  $L_c$ . (b) Comparison between  $L_s^*(L_c)$  (circles) and the square-root of the mean contributing area,  $\langle a(b) \rangle^{1/2}$  (line), where both  $L_c$  and b are averaging lengths signified by x. (c) Histogram showing the frequency of occurrence of  $\delta\theta$  for the case  $L_c = 1.0$  km and  $L_s = L_s^*$ ; i.e., the maximum correlation in (a).

the case  $\lambda_0 = 0.46$  km. The results should, in general, depend on  $\lambda_0$ , but this dependence is beyond the scope of the present study.) Although the correlations are never large, one striking feature stands out: the maximum occurs when  $L_s \approx 10L_c$ , where here  $L_c = 1.0$  km. In other words, the direction in which channel-like features flow locally is most strongly coupled to the direction of the topographic slope measured at a *much larger scale*. We call this large scale  $L_s^*$ . For any choice of  $L_c$ , a curve similar to that shown in Fig. 1(a) is obtained, but with its maximum at a different  $L_s$ . Thus,  $L_s^*$  is a function of  $L_c$ .

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We next provide empirical evidence that  $L_s^*$  represents the average size of the individual basins that drain into a particular site on a map. Using the usual procedure [2], we calculate the area  $a_i$  that drains into the *i*th location. We calculate  $\{a_i\}$  not only on the original map but also on maps coarse grained at a scale *b*. (Coarse graining is performed by computing the mean elevation in blocks of linear dimension *b*.) The average  $\langle a(b) \rangle$  then represents the mean contributing area for the coarse-grained map, and a characteristic length scale may be obtained from its square root. Because both *b* and  $L_c$  are averaging length scales, we may compare  $\langle a(b) \rangle^{1/2}$  to  $L_s^*(L_c)$  for  $b = L_c$ . Figure 1(b) shows that the quantitative agreement is surprisingly good.

Another view of the two-scale interaction is shown in Fig. 1(c), a histogram of  $\delta\theta$ , for the case  $L_s = 10$  km and  $L_c = 1$  km, i.e., the point that gives the maximum in Fig. 1(a). One sees that the maximum probability is for  $\delta\theta \approx 0$ . In other words, the nonlocal coupling of the scales  $L_s$  and  $L_c$  results in a tendency, on average, for the channelization vector **c** and the slope vector **s** to be parallel, but only when  $L_s \gg L_c$ . Note that isotropic random topography would yield a flat histogram for any choice of  $L_s$  and  $L_c$ . Thus, the results of Fig. 1 may have some practical use in the identification of the effects of fluvial erosion in environments, such as Mars [25], where the origin of channel-like features is unclear.

We may provide a more detailed view of the two-scale interaction by explicitly incorporating the drainage area a in our correlations. We include the *m*th moment of a in our measure of the locally averaged slope by defining

$$\widetilde{\mathbf{s}}_{m}(\mathbf{r};L_{s}) = G^{-1} \int_{|\mathbf{x}-\mathbf{r}| \le L_{s}/2} d\mathbf{x} \ g(\mathbf{x}-\mathbf{r};L_{s}) a^{m}(\mathbf{x}) \cdot \boldsymbol{\nabla} h(\mathbf{x}).$$
(8)

We then compute the scale-dependent average

$$W_m = \frac{\langle |\mathbf{c} \cdot \widetilde{\mathbf{s}}_m|^2 \rangle}{\langle |\mathbf{c}|^2 | \widetilde{\mathbf{s}}_m|^2 \rangle},\tag{9}$$

where  $\tilde{\mathbf{s}}_m$  and  $\mathbf{c}$  are measured at scales  $L_s$  and  $L_c$ , respectively. The parameter *m* controls the statistical weight given to regions with large contributing area; the higher the moment *m*, the greater the weight given to slopes with large contributing area.

Figure 2(a) shows how  $W_m$  depends on  $L_s$ ,  $L_c$ , and m. Three trends are worthy of note: (i) the maximum  $W_m^*$  of  $W_m$  increases with increasing m [Fig. 2(a)]; (ii)  $W_m^*$  occurs at a value  $L_s^*$  that decreases with m [Figs. 2(a) and 2(b)]; and (iii)  $W_m(L_s)$  is insensitive to  $L_c$  [Fig. 2(c)].

The first trend shows that, on average, **c** is most correlated with **s** when the drainage area is large. In other words, the two-scale interaction is strongest for channels that drain large basins. Recall additionally that  $a(\mathbf{r})$  is a highly fluctuating function of space [2] (i.e., its variance is much larger than its mean), whereas **c** is not. The locations that contribute most to  $W_m$  are therefore where *a* is large. These locations become increasingly localized (i.e., they are less numerous, less dis-



FIG. 2. (a) The correlation function  $W_m$  for the case  $L_c = 5.7$  km, for m=1 (circles), m=3 (squares), and m=5 (triangles). (b) Plot of  $\log_{10} L_s^*$  as a function of the moment  $\log_{10} m$  (circles), where  $L_s^*$  is the value of  $L_s$  that maximizes  $W_m$ . The data are compared with a straight line of slope -0.3, indicating that  $L_s^*$  scales approximately as  $m^{-0.3}$ . (c)  $W_m$  as a function of  $L_s$  for  $L_c = 2.9$  (circles), 4.7 (squares), and 6.6 km (triangles) for the case m=3. A similar insensitivity of the shape and position of the curves with respect to  $L_c$  is found for other m.

tributed, and stronger) as *m* increases. Therefore the averaging scale  $L_s$  over which  $\tilde{\mathbf{s}}_m$  correlates with **c** decreases with *m*, thus explaining the second trend. The third trend results from the slow variation of **c** compared to  $\tilde{\mathbf{s}}_m$ .

These results may provide some insight into phenomenological continuum-mechanical theories of erosion. Typically, these formulations assume, directly or indirectly, that the flux **j** of eroded material is proportional [26] to a product of drainage area and slope [5-16], i.e.,

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where  $s = |\mathbf{s}|$ ,  $\hat{\mathbf{s}} = \mathbf{s}/s$ , and *n* is another parameter. Our term  $\tilde{\mathbf{s}}_m$  corresponds precisely to the right-hand side of Eq. (10) with n=1, averaged over a scale  $L_s$ . Now assume that the left-hand side, when averaged over a scale  $L_c$ , may be approximated to first order by the channelization vector c. (This assumption may be partially justified by noting that deeper channels carry more sediment.) In other words, we view  $\tilde{\mathbf{s}}_m$  as a generalized force and  $\mathbf{c}$  as an approximate flux. For this interpretation to be correct,  $\tilde{s}_m$  and c should be defined at the same scale L, and the maximum correlation  $W_m$ should occur at this scale. In Fig. 2(a) these two conditions are satisfied by the curve  $W_3$ , which reaches its maximum at  $L_s = L_c = 5.7$  km. Thus, when the averaging scale L = 5.7 km, the "best" effective equation of form (10) with n=1 appears to require m=3. More generally, since Fig. 2(c) shows that  $W_m$  is independent of  $L_c$ , we may read directly from Fig. 2(b) that  $L \simeq L_0 m^{-\alpha}$ , with  $L_0 \simeq 0.9$  km and  $\alpha \simeq 0.3$ , giving, upon inversion,  $m \simeq (L/L_0)^{-1/\alpha}$ .

The main conclusion to be derived from Fig. 2 lies not with these specific numbers or relations but with the overall trend: small m characterizes large length scales, while large m characterizes small length scales. These observations are consistent with the notion that a diffusion equation [27] (corresponding to m=0, n=1) is a zeroth-order model of erosion. Higher powers of a in Eq. (10) then correspond to higher-order corrections that may be identified with smaller-scale features. Note, however, that our analysis provides no indication that these higher-order corrections become progressively smaller. Indeed, the trends in Fig. 2(a) point in much the opposite direction.

Our results do not necessarily invalidate Eq. (10) as an effective transport law. Its use is often motivated by the observation that  $\langle a^m s^n \rangle \simeq \text{const}$  for an appropriate choice of m/n [2]. However, Ref. [28] shows that  $\langle a^m s^n \rangle \simeq \text{const}$  also holds for random topography (i.e., Gaussian surfaces). Since random topography implies random **j**, Eq. (10) cannot be valid for a random surface. Here, on the other hand, we find it invalid only in a scale-independent sense.

In conclusion, we have presented a multiscale analysis of an eroding landscape that explicitly quantifies the coupling of drainage basins to channels. This coupling has been shown to create a two-scale interaction, which is itself scale dependent. These results indicate that a commonly employed effective equation for erosive transport also contains hidden dependencies on scale. While our results indicate that the construction of a comprehensive continuum theory for erosion is a formidable challenge, they do show one way in which landscape patterns of unknown origin [25] may be quantitatively analyzed to determine the kind of mechanisms that have eroded them. We hope that our methods may also find some applicability in the analysis of other types of landscapes, such as those formed by fracture [29], where the underlying dynamics are poorly understood.

We thank P. Dodds, R. Pastor-Satorras, A. Rinaldo, and N. Schorghofer for helpful discussions. This work was supported in part by DOE Grant No. DE FG02-99ER 15004 and NSF Grant No. EAR-9706220.

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